

1                   **Ignoring internal variability can lead to misleading conclusions on model**  
2                   **fidelity**

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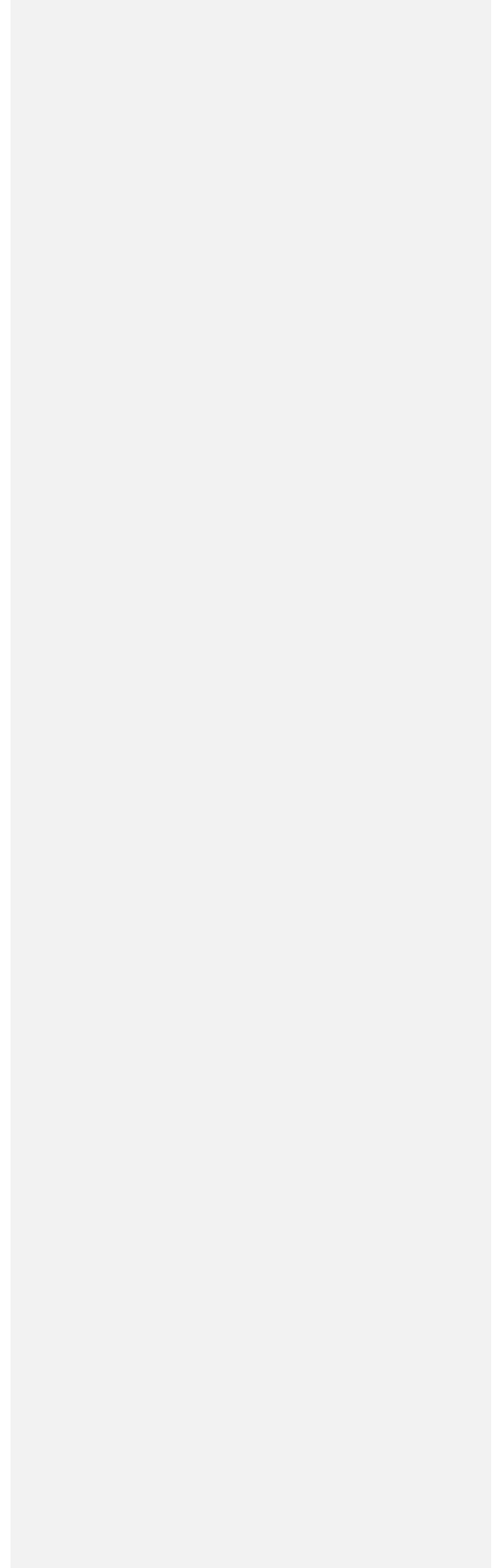
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13 **Abstract**

14 Benchmarking climate model simulations against observations of the climate is core to the  
15 process of building realistic climate models and developing accurate future projections.  
16 However, in many cases, models do not match historical observations, particularly on regional  
17 scales. If there is a mismatch between modeled and observed climate features, should we  
18 necessarily conclude that our models are deficient? Using several illustrative examples, we  
19 emphasize that internal variability can easily lead to marked differences between the basic  
20 features of the model and observed climate, even when decades of model and observed data  
21 are available. This can appear as an apparent failure of models to capture regional trends or  
22 changes in global teleconnections, or simulation of extreme events. Despite a large body of  
23 literature on the impact of internal variability on climate, this acknowledgment has not yet  
24 penetrated many model evaluation activities, particularly for regional climate. We emphasize  
25 that using a single or small ensemble of simulations to conclude that a climate model is in  
26 error can lead to premature conclusions on model fidelity. A large ensemble of multidecadal  
27 simulations is therefore needed to properly sample internal climate variability in order to  
28 robustly identify model deficiencies and convincingly demonstrate progress between  
29 generations of climate models.

30 **Keywords:** Chaotic internal variability, climate model evaluation, future projections, large  
31 ensembles.

## 32 1. Introduction

33 Climate models are remarkably successful in reproducing many earth-system  
34 phenomena such as atmospheric jet streams, oceanic currents, monsoons, the El Niño  
35 Southern Oscillation (ENSO), the North Atlantic Oscillation (NAO), and the response of global  
36 climate to external forcing (Broecker, 1975; Hansen et al., 1981, Hausfather et al., 2020). From  
37 their basis in the Navier-Stokes equations of fluid dynamics, even extreme events like  
38 heatwaves and cold snaps, floods and droughts, cyclones, and storms all appear  
39 spontaneously in climate model simulations. In some cases, models also warn us of more  
40 intense extreme events than we have yet experienced but which could plausibly occur at any  
41 time in the current climate (Jain et al., 2020; McKinnon and Deser, 2021; Thompson et al.,  
42 2017; Kent et al., 2022). Scientists also use climate models to understand the physical  
43 mechanisms behind past changes in climate (e.g., Mitchell et al., 1995), to understand and  
44 predict extreme events (e.g., Hardiman et al., 2020), to project future climate change, and to  
45 inform governments and policymakers about the impacts of climate change (IPCC AR6 WG I  
46 report, 2021).

47 To gain confidence in model projections, the fidelity of models is assessed by  
48 comparing model simulations of the historical climate with observations. This trial of models  
49 using observations is core to identifying current model deficiencies and prioritizing areas for  
50 further development (e.g., Flato et al., 2013). The scientific literature presents many examples  
51 where climate models apparently fail to reproduce even basic aspects of observed regional  
52 climate such as trends in regional rainfall amount and temperature (e.g., Mishra et al., 2018;  
53 Mitchell et al., 2013; Saha et al., 2014), multidecadal changes in atmospheric circulation  
54 (Gillett et al., 2003) and climatology (Seidel et al., 2008), the frequency or magnitude of  
55 extreme events (Tyrrell et al., 2022, Slingo et al., 2022), global teleconnections (Butler and  
56 Polvani, 2011), interaction between different modes of climate variability (Martin et al., 2021)  
57 or external forcing effects (Schmidt et al., 2014; Legrande et al., 2016). However, even in the  
58 case of perfect models, perfect boundary conditions, and perfect observations, a lack of

59 agreement between the modeled and observed climate can still arise simply due to chaotic  
60 internal variability (Tebaldi and Knutti, 2007; Deser et al., 2012, 2016, 2018; Mitchell et al.,  
61 2013; Jain and Scaife, 2022; Fasullo et al., 2020, McKinnon and Simpson, 2022).

62 Due to the inevitable presence of internal variability, each realization of climate, in both  
63 observed and models, represents only a possible state (Lorenz 1963, 1969). When evaluating  
64 models against observations, this issue (known as sampling variability) is supposed to be  
65 handled through statistical tests, but the power of those tests relies heavily on the assumed  
66 chance process generating internal variability. This leads to a misinterpretation of significance  
67 tests, which is widespread in climate science, leading to overly lax criteria (Shepherd, 2021).  
68 This problem is only exacerbated if the statistical tests do not take proper account of  
69 multidecadal internal variability, which is very difficult to quantify from single or small model  
70 ensembles of models, or from observations.

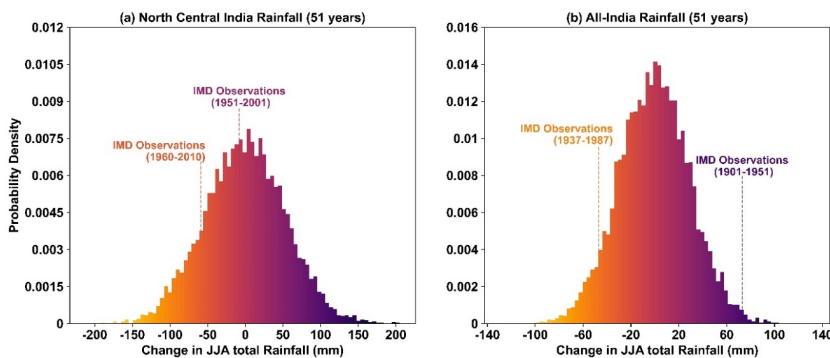
71 Many of the studies claiming a mismatch between models and observations, including  
72 some cited in the IPCC WG1 report (2021), continue to use only a single or a small ensemble  
73 of simulations from a given model. These studies generally show that the differences are  
74 *statistically significant*, typically for time period considered sufficiently long. Here, we provide  
75 three illustrative examples to demonstrate how internal variability cannot be easily ruled out  
76 as a cause of commonly reported discrepancies even when decades of observations and  
77 model data are available.

## 78 **2. Do models reproduce observed trends?**

79 Over the latter half of the 20<sup>th</sup> century, historical observations show a significant  
80 reduction in the summer monsoon rainfall over parts of India (Kumar et al., 2010, Ramesh et  
81 al., 2014; Saha et al., 2014; Jin and Wang, 2017). However, most historical climate model  
82 simulations from different CMIP generations have shown a consistent increase in rainfall  
83 during this period and beyond into the 21st century under increasing greenhouse gas forcing  
84 (Katzenberger et al., 2021; Saha et al., 2014; Shashikanth et al., 2013).

85 Several hypotheses have been proposed to explain the apparent mismatch in  
86 observed and model-simulated trends in monsoon rainfall over north-central India, including  
87 large observational uncertainty (Singh et al., 2019), the recent warming of the Indian Ocean  
88 (Roxy et al., 2015), radiative forcing due to aerosols (Monerie et al., 2022) or changes in land-  
89 use and land-cover (Paul et al., 2016). In most cases, these factors are found to have  
90 systematic effects on the modeled monsoon rainfall. However, it is also possible that internal  
91 variability on multidecadal timescales contributes to the mismatch.

92 We find that the most extreme drying and wetting trends in the Indian Meteorological  
93 Department (IMD) observational record for both regional (Fig. 1a) as well as all-India rainfall  
94 (Fig. 1b) lie within the range of plausible trends from chaotic internal variability in the current  
95 climate. Even more extreme values than have been observed in the IMD record so far are  
96 possible, solely due to internal variability, in the absence of any systematic forced effects.  
97 Therefore, for this case, internal variability cannot easily be rejected as the cause of the  
98 models' apparent failure to capture the observed trend (also see Huang et al., 2020; Sinha et  
99 al., 2015).



100

101 **Figure 1 Internal variability in rainfall trends.** Change in June-July-August (JJA) total  
102 rainfall over a 51-year period from internal variability in a multimodel ensemble (MME) of  
103 climate predictions for (a) north-central India (20-28°N, 76-87°E) and (b) all-India (land-only)  
104 rainfall. The darker color indicates wetter trends. The dotted lines show the most extreme 51-  
105 year trends in the IMD observational record during 1950-2013 (the period for which drying  
106 trends were observed over north-central India) and 1901-2013 for all-India. See Methods for  
107 more details.

108 Unprecedented climate extremes are often a manifestation of both internal variability  
109 and external forcing. However, in many cases, the internal variability is so large that it can  
110 easily negate or greatly overwhelm any forced response in climate trends, even on  
111 multidecadal time scales (e.g. Fischer et al., 2013; Jain and Scaife, 2022). For example, in  
112 this case, internal variability can be large enough to overwhelm the wetting trend due to  
113 greenhouse gas forcing and give temporary drying trends in monsoon rainfall.

114 Another example where the role of internal variability was ignored is the recent paper  
115 by Scafetta (2022), which claimed that no model with a climate sensitivity  $>3^{\circ}\text{C}$  was consistent  
116 with observed trends since 1979. However, this claim is based on comparing each model's  
117 ensemble mean change with an observational estimate (from ERA5) without taking either  
118 observational uncertainty or model internal variability into account. When these elements are  
119 included and appropriate statistical tests were employed (following Santer et al, 2008), it was  
120 demonstrated that the majority of the models with sufficiently large ensembles, even some of  
121 those with high climate sensitivity of  $5^{\circ}\text{C}$ , are consistent with the observations (Schmidt et al.  
122 2023).

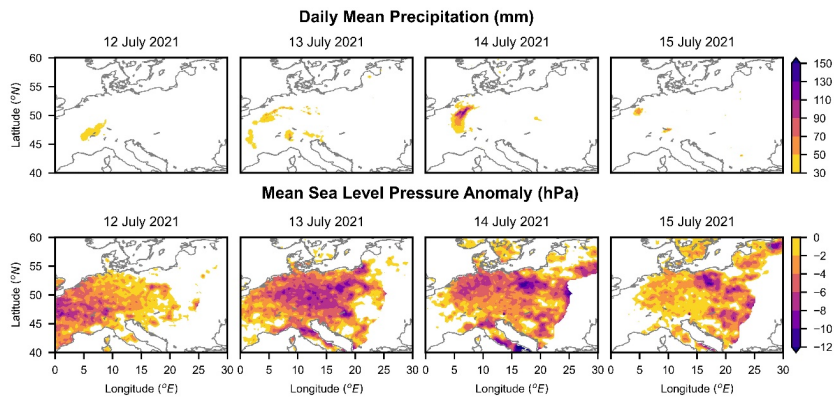
123 There also exists a selection bias when comparing the most extreme trends in the  
124 observational record with model-simulated trends (Rahmstorf et al. 2017). Inspection of  
125 observational time series, with no a priori reason to select the particular period of most extreme  
126 trends, followed by comparison with model-simulated trends *for the same time period*  
127 introduces a selection bias and the impression that models fail to produce observed trends. In  
128 many cases, those extreme trends in regional climate can appear at any time in the model  
129 simulation and not necessarily during the observed period, irrespective of any forced changes  
130 which are often smaller than the internal variability for short- or medium-term periods. It is  
131 therefore very difficult to argue that models cannot reproduce observed extremes by pre-  
132 selecting extreme periods in the observations and then testing models for the same historical  
133 period. Significance testing of these extreme trends is difficult in such cases and simple tests

134 based on a single or limited ensembles of simulations to reject the null hypothesis of internal  
135 variability in favor of a model error are often invalid (Eade et al., 2021).

136 Whilst the Indian rainfall trend presented here is only one example, many other cases  
137 of models apparently failing to reproduce observed trends for other regions and variables also  
138 exist (e.g. Gillett et al., 2003, Seidel et al., 2008, Swart et al. 2015, Deser et al., 2020, 2023).  
139 Therefore, it is necessary to re-examine such cases and carefully rule out internal variability  
140 as a cause of apparent mismatch between observed and modeled trends to robustly identify  
141 the true model errors.

### 142 **3. How should we test models for extreme events?**

143 In addition to the selection bias in time, there also exists a selection bias in space. If  
144 we preselect a particular extreme event from the observational record, which, by definition, is  
145 a rare event, and look for similar events in climate model simulations, then the chances of  
146 finding rare events of the same magnitude, duration, and spatial scale, at the same location  
147 will necessarily be low. However, in this case, it is premature to then conclude that models  
148 cannot simulate the observed extreme event. For instance, if we then use a large ensemble  
149 of climate simulations and search for a similar intense event with *no a priori specification of*  
150 *exactly where it should happen*, it is often possible to find a similar extreme event in the  
151 simulations, leading to very different conclusions about model fidelity (e.g., Fischer et al. 2013,  
152 2021).



153

154 **Figure 2a. Observed extreme rainfall over Europe in July 2021.** Daily mean rainfall (mm)  
 155 over Europe for 12-15 July (top panels) and corresponding daily mean sea level pressure  
 156 anomaly (hPa) (bottom panels) from E-OBS v26.0 dataset. The daily sea level anomalies are  
 157 calculated with respect to monthly values for July 2021.

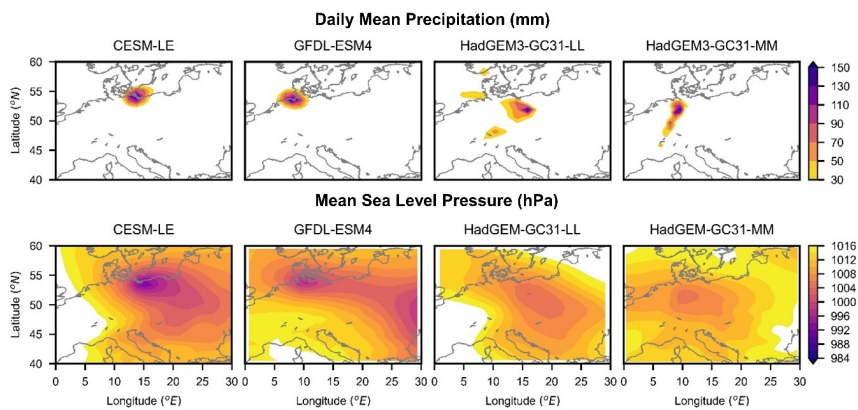
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159 To illustrate this, we use the example of the German floods of 2021. The observed  
 160 event in July 2021 was associated with daily mean rainfall reaching as high as 150 mm over  
 161 parts of Germany (Fig. 2a). Searching ensembles of climate simulations from multiple models  
 162 for European rainfall extremes of similar magnitude to that observed, reveals several instances  
 163 with rainfall intensity in climate models reaching, and in some cases even exceeding the  
 164 rainfall rate seen in the observational record (Fig. 2b). We can also examine the mean sea  
 165 level pressure (MSLP) to determine if the simulated extreme corresponds to a realistic  
 166 circulation pattern. In all cases, the extreme rainfall events in the models are co-located with  
 167 extreme low-pressure regions, indicating low-level convergence and enhanced probability of  
 168 intense rainfall, similar to the observed event which had low MSLP in the vicinity.

169 To illustrate the point further, we also examined a large ensemble from a single model,  
 170 the CESM1 (Kay et al., 2015) as this isolates the impact of internal variability. While we find  
 171 similar mid-latitude extremes occurring randomly anywhere over the selected domain in some  
 172 simulations of the large ensemble, other simulations from the same model did not simulate  
 173 any such events. This shows that a-priori constraining the regional scale and location in the



174 model to that of the pre-selected local record event invalidates commonly used statistical  
175 significance tests and hence can lead to the erroneous conclusion that models cannot  
176 reproduce the observed extreme.



177

178 **Figure 2b. Simulation of extreme European rainfall in climate models.** Daily mean rainfall  
179 (mm) over Europe for a selected day between 1950-2014 (top panels). Corresponding daily  
180 mean sea level pressure (hPa) is shown in the bottom panels. The CESM-LE case is from the  
181 CESM1 Large Ensemble and the remaining models are from CMIP6 ensembles.

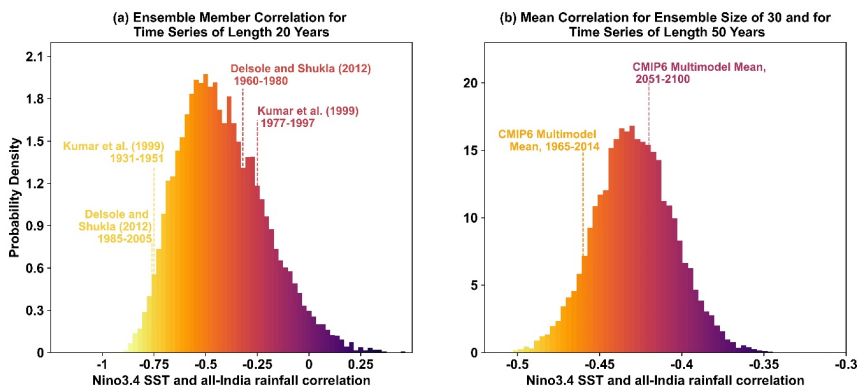
#### 182 4. Are teleconnections changing?

183 Large-scale internal variability from phenomena such as ENSO, NAO, or the Indian  
184 Ocean Dipole (IOD), are known to influence regional climate across the globe through  
185 *teleconnections*. These teleconnections lead to systematic changes in regional climate far  
186 from the center of action of the variability itself (e.g., Branstator, 2002; Hurrell et al., 2003;  
187 Trenberth et al., 2020; Yamagata et al., 2004) and often contribute to the predictability of  
188 regional weather and climate events. For example, the ENSO-rainfall teleconnection is crucial  
189 for tropical rainfall prediction at seasonal lead times (Jain et al., 2019, Scaife et al., 2019).

190 Recent literature questions the stationarity of these and other teleconnections and in  
191 many cases argues that causal mechanisms, such as externally forced climate change, are  
192 driving systematic change in either the pattern or the strength of the teleconnection from one  
193 epoch to another, and that these changes are not represented in climate models. For instance,  
194 changes in the ENSO-rainfall teleconnection have been reported for several tropical regions

195 including India (e.g. Krishna Kumar et al., 1999), East Asia (e.g. Wu and Wang 2002), North  
 196 America (e.g. Coats et al., 2013), and Africa (e.g. Losada et al., 2012, Bahaga et al., 2019),  
 197 as well as many other large-scale teleconnections, such as the recently discovered connection  
 198 between the Quasi-Biennial Oscillation and Madden Julian Oscillation (Lee and Klingaman,  
 199 2018; Martin et al, 2021) and the ENSO-Asian teleconnection (He et al., 2013).

200 Several studies already highlight internal variability as a cause of the apparent  
 201 mismatch between model and observed teleconnections (Kitoh 1997, Bodai et al., 2020, Deser  
 202 et al., 2017; Lee and Bódai, 2021, Yun and Timmerman, 2018) but many others continue to  
 203 suggest that mismatch implies model error. Therefore, we re-examine one well-known  
 204 example: ENSO and Indian summer monsoon rainfall tele-connection.



205  
 206 **Figure 3 Internal variability in the ENSO-monsoon relationship.** Probability distribution of  
 207 Pearson's correlation coefficient between detrended Nino3.4 sea-surface temperatures  
 208 (SSTs) and rainfall from the CHFP MME. Panel (a) is for resampled rainfall and SST time  
 209 series of length 20 years and shows the range of correlations corresponding to individual  
 210 ensemble member time series. Panel (b) is for 50-year time series and correlation values were  
 211 randomly resampled in sets of 30 and therefore represent the range of *mean* correlations for  
 212 30 ensemble members (see Methods). The dotted lines show some extreme examples of  
 213 correlation values from the literature for corresponding timescales. Examples shown in panel  
 214 (b) are from Lee and Bodai (2021) for the ensemble size of 30.

215 Figure 3 (a) shows that the distribution of ENSO teleconnections in rainfall resamples  
 216 covers an enormous range of correlations on 20-year timescales ( $r = -0.90$  to  $0.47$ ). A similar  
 217 result holds for 50-year timescales ( $r = -0.80$  to  $0.22$ ). Note that this range occurs due to

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219 sampling variability rather than any true systematic non-stationarity, and includes the historical  
220 periods such as 1980-2010 when the non-stationarity in the ENSO-Indian rainfall  
221 teleconnection has been reported in observations (e.g. Krishna Kumar et al., 1999). Whilst  
222 methodologies to calculate the ENSO-monsoon relationship vary in the literature, even the  
223 extreme examples of apparent non-stationarity in observed teleconnections sit well within the  
224 spread of plausible ENSO-monsoon teleconnections due to the internal variability of the  
225 climate system. The model resamples even show the possibility of a positive correlation on  
226 20-year timescales; opposite in sign to the observed teleconnection.

227 In addition, there is also a growing body of literature suggesting that the ENSO-Indian  
228 rainfall teleconnection could change in the future under climate change (e.g. IPCC WG1  
229 report, Chapter 14, 2013; Roy et al., 2019). However, Fig. 3 (b) shows that the mean ENSO-  
230 rainfall correlation for the CMIP6 multimodel ensembles (of size  $\sim 30$ ) for both historical and  
231 future periods (Lee and Bodai, 2021) sits within the range of internal variability ( $r = -0.31$  to -  
232 0.51). Therefore, great care is needed before we can conclude with confidence that there is  
233 any robust change in the ENSO-monsoon relationship in the future.

234 Finally, given the broad range of possibilities in Fig. 3, and the fact that simulated future  
235 changes are well within this range, it is unlikely that observational data will yield significant  
236 examples of changes in teleconnections that are extreme enough to rule out internal variability  
237 and detect any true non-stationarity on any reasonable timescale into the future.

## 238 **5. A call for more rigorous model assessment**

239 We find that studies claiming a mismatch between model and observed climate  
240 phenomena are often too quick to ignore the null hypothesis that such apparent discrepancies  
241 between models and observations can arise due to chaotic internal variability. We have  
242 presented examples where this applies to changes on multidecadal timescales, such as global  
243 and regional trends, recent extreme events, and apparent changes in observed  
244 teleconnections.

245           Assessing models against the rare and most extreme observed cases automatically  
246 introduces a selection bias into the process of model evaluation which is particularly  
247 compounded for extreme events on small spatial scales. Limited observational records can  
248 easily show apparent non-stationary or spurious teleconnections due to internal variability and  
249 sampling error. Using a single simulation or small ensembles of simulations can easily lead to  
250 premature conclusions about model fidelity. Therefore, a large ensemble of multidecadal  
251 simulations is needed to properly sample initial condition uncertainty and demonstrate model  
252 failure to capture observed phenomena, or the progress or deterioration of performance  
253 between old and new generations of models, including costly new climate models at high  
254 resolution. Initialized large ensembles are already being used for seasonal predictions  
255 (Buontempo et al., 2018; Tompkins et al., 2017), and more recently for multidecadal  
256 predictions (Boer et al., 2016) and projections (Deser et al., 2020, Maher et al., 2020). Using  
257 these ensembles to isolate internal variability from true model errors provides a powerful  
258 second application.

## 259 **6. What else can be done?**

260           While a large ensemble is desirable to account for fluctuations due to internal  
261 variability, we acknowledge that these are computationally expensive and may not be always  
262 possible. Therefore, we also highlight other potential model evaluation methodologies that  
263 could be used. For instance, in contrast to picking a single period in observations and testing  
264 a single model simulation against that, we suggest using a longer time period, and sampling  
265 all possible periods of a fixed length within that interval. For example, sampling 20-year trends  
266 over 50 years for both observations and models and comparing the distribution of trends (cf.  
267 Jain and Scaife 2022, Scaife et al. 2009, Eade et al. 2022). Comparing multimodel mean  
268 trends directly with the observations is also not a fair comparison (Schmidt et al. 2023) and for  
269 these cases, statistical tests, similar to the UNprecedented Simulated Extremes in Ensembles  
270 (UNSEEN) method (Thompson et al., 2017), could be employed to test model fidelity.

271 Grid point comparisons for extremes, such as calculating spatial distributions of  
272 extremes (e.g., 1-day maximum rainfall) and comparing those with the model spatial  
273 distribution, are likely to show apparent disagreement. A similar problem has been recognized  
274 as 'double-penalty' in high-resolution weather prediction where verification scores are  
275 penalized twice, i.e. for simulating a feature in the wrong place and not simulating a feature at  
276 the right place (Hagelin et al., 2014). Simple approaches, such as pooling daily maximum  
277 rainfall values over a spatial domain and time period for both observations and models and  
278 comparing those distributions, or more sophisticated methodologies (such as Hi-RA,  
279 Mittermaier, 2014) could be used for evaluating climate models for extremes.

280 There are also new emerging methodologies, such as resampling of observational  
281 records to create pseudo observational large ensembles (McKinnon and Deser, 2018, 2021),  
282 quantification of the degree of discrepancy between models and observations using Bayes  
283 factors (Shepherd, 2021), spatial aggregation for robust projection of regional extremes using  
284 current models (Fischer et al. 2013), or using physical or process-based evaluation such as  
285 storyline methods (Shepherd et al., 2018) for post-event simulation of extreme observed  
286 events in models, that could be employed for better control on internal variability and sampling  
287 error. Finally, we highlight that models will always contain errors to some extent. However, to  
288 robustly determine true errors in climate models and to make an informed statement about  
289 them, it is important to carefully acknowledge the role of internal variability.

## 290 **Methods**

291 **Figure 1** The ensemble member rainfall was bias-adjusted using the difference between the  
292 ensemble mean and observations for each model's hindcast period. Each model's fidelity was  
293 tested using the UNSEEN method and rainfall MME was created using a selected set of  
294 models that passed the UNSEEN tests (GloSea5, CFS, and the MPI models for north-central  
295 India for 1950-2013, and GloSea5, ECMWF-S4, CFS, MIROC5, and MPI-LR for all-India for  
296 1901-2013) (see Table 1 of Jain et al, 2020 for details of models, Tompkins et al., 2017). The  
297 ensemble mean trend was removed from each ensemble member time series for each model

298 to remove any influence of climate change before creating MME. A total of 10,000 time series  
299 of length 51 years were resampled from MME by randomly selecting ensemble members in  
300 sets of three consecutive years to preserve the interannual autocorrelation in rainfall. Linear  
301 trends were calculated as the slope of the line of best fit for each time series to obtain 10,000  
302 values of trends. Each trend value was then multiplied by the length of the time series (i.e. 51  
303 years) to obtain the change in JJA total rainfall over a 51-year period, shown on the x-axis.  
304 We also tested the sensitivity of the extreme values shown in Fig. 1 to the model variance by  
305 removing one of the models with the highest variance while resampling. We find no substantial  
306 influence of the differences in the models' variance on the extreme values presented in Fig. 1.

307 **Figure 3** The rainfall MME described in Fig. 1 was used to randomly resample 10,000 time  
308 series of rainfall and corresponding sea-surface temperatures (SSTs) of lengths 20 and 50  
309 years. The correlation coefficient was calculated for each of the 10,000 time series for both 20  
310 (Fig. 3a) and 50 years (Figure not shown). The SSTs were averaged over the Nino3.4 region  
311 (5°S to 5°N, 170°W to 120°W) and detrended before sampling to remove the influence of  
312 forced climate change on SSTs. For Fig. 3b, we randomly selected 30 correlation values from  
313 the 10,000 correlation values for the length of 50 years. We then calculated the average of the  
314 selected 30 correlation values to obtain the *mean* correlation and this was repeated 10,000  
315 times to obtain the distribution in Fig. 3b.

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560 [cmip6](https://cds.climate.copernicus.eu/cdsapp#!/dataset/projections-cmip6), <https://www.wcrp-climate.org/wgsip-chfp/chfp-data-archive> and  
561 <https://www.cesm.ucar.edu/projects/community-projects/LENS/data-sets.html>. The IMD  
562 rainfall data can be obtained from the following  
563 link: [https://www.imdpune.gov.in/Clim\\_Pred\\_LRF\\_New/Grided\\_Data\\_Download.html](https://www.imdpune.gov.in/Clim_Pred_LRF_New/Grided_Data_Download.html).

564 The authors declare no conflict of interest.